

Journal Pre-proof

Association of obstructive sleep apnea with severity of patients hospitalized for acute asthma

Shojiro Oka, MD, Tadahiro Goto, MD, MPH, Atsushi Hirayama, MD, MPH, Mohammad Kamal Faridi, MPH, Carlos A. Camargo, Jr., MD, DrPH, Kohei Hasegawa, MD, MPH

PII: S1081-1206(19)31382-1

DOI: <https://doi.org/10.1016/j.anai.2019.11.002>

Reference: ANAI 3062

To appear in: *Annals of Allergy, Asthma and Immunology*

Received Date: 18 August 2019

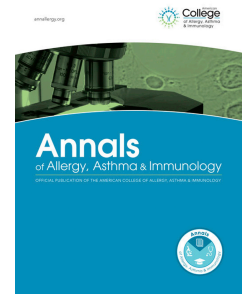
Revised Date: 24 October 2019

Accepted Date: 5 November 2019

Please cite this article as: Oka S, Goto T, Hirayama A, Faridi MK, Camargo CA Jr., Hasegawa K, Association of obstructive sleep apnea with severity of patients hospitalized for acute asthma, *Annals of Allergy, Asthma and Immunology* (2019), doi: <https://doi.org/10.1016/j.anai.2019.11.002>.

This is a PDF file of an article that has undergone enhancements after acceptance, such as the addition of a cover page and metadata, and formatting for readability, but it is not yet the definitive version of record. This version will undergo additional copyediting, typesetting and review before it is published in its final form, but we are providing this version to give early visibility of the article. Please note that, during the production process, errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

© 2019 American College of Allergy, Asthma & Immunology. Published by Elsevier Inc. All rights reserved.



1 **Association of obstructive sleep apnea with severity of patients hospitalized for acute**
2 **asthma**

3
4 Shojiro Oka, MD¹; Tadahiro Goto, MD, MPH^{2,3,4}; Atsushi Hirayama, MD, MPH^{2,5}; Mohammad
5 Kamal Faridi², MPH; Carlos A. Camargo, Jr., MD, DrPH²; and Kohei Hasegawa, MD, MPH²

6
7 **Affiliations:**

- 8 1. Department of Emergency Medicine, Okinawa Prefectural Chubu Hospital, Okinawa, Japan
9 2. Department of Emergency Medicine, Massachusetts General Hospital, Boston, MA, USA
10 3. Graduate School of Medical Sciences, University of Fukui, Fukui, Japan
11 4. Department of Clinical Epidemiology and Health Economics, School of Public Health, The
12 University of Tokyo, Tokyo, Japan
13 5. Department of Cardiology, National Cerebral and Cardiovascular Center, Osaka, Japan

14
15 **Email addresses:**

16 Shojiro Oka: oka_shojiro@hosp.pref.okinawa.jp
17 Tadahiro Goto: tag695@mail.harvard.edu
18 Atsushi Hirayama: ath877@mail.harvard.edu
19 Mohammad Kamal Faridi: mkfaridi1@mgh.harvard.edu
20 Carlos A. Camargo, Jr.: CCAMARGO@PARTNERS.ORG
21 Kohei Hasegawa: KHASEGAWA1@PARTNERS.ORG

22

23 **Corresponding Author:** Dr. Tadahiro Goto, Graduate School of Medical Sciences, University of
24 Fukui, Matsuoka-shimoaizuki, Eiheiji, Yoshida, Fukui, 910-1193, Japan; and Department of
25 Emergency Medicine, Massachusetts General Hospital, Boston, MA, 02114-1101, USA. Email:
26 tag695@mail.harvard.edu

27

28 **Author contributions:** T.G. takes responsibility for the paper as a whole. S.O., T.G., A.H., and
29 K.H. conceived the study. C.A.C. obtained research funding. T.G., C.A.C. and K.H. supervised
30 the conduct of the study. T.G., A.H., M.K.F., C.A.C. and K.H. provided statistical advice. T.G.,
31 A.H., and M.K.F. analyzed the data. S.O. and T.G. drafted the manuscript, and all authors
32 contributed substantially to its revision.

33

34 **Conflict of Interest:** Dr. Camargo has provided asthma-related consulting services to
35 AstraZeneca and GlaxoSmithKline. Dr. Hasegawa has received grants for asthma-related
36 research from Novartis and Teva. The other authors have no relevant financial relationships to
37 disclose.

38

39 **Funding:** This study was supported by the grant R01 HS-023305 (Camargo) from the Agency
40 for Healthcare Research and Quality (Rockville, MD). The content of this manuscript is solely
41 the responsibility of the authors and does not necessarily represent the official views of the
42 Agency for Healthcare Research and Quality.

43

44 **Running head:** Obstructive sleep apnea and acute asthma severity

45

46 **Notation of prior abstract publication/presentation:** This study was presented at the American
47 Thoracic Society International Conference, San Diego, CA, USA; May 2018

48

49 **Word count:** 2,469 words; 2 Tables; and 1 Figure

50

51 **Abbreviations:**

52 CI, confidence interval

53 HCUP, Healthcare Cost and Utilization Project

54 *ICD-9-CM, International Classification of Diseases, Ninth Revision, Clinical Modification*

55 LOS, length-of-stay

56 OR, odds ratio

57 OSA, Obstructive sleep apnea

58 SID, State Inpatient Databases

59

60 **Key words:** Obstructive sleep apnea; acute asthma; hospitalization; severity; positive pressure
61 ventilation; length-of-stay

62

19-08-0408R2

Background: Studies suggest that obstructive sleep apnea (OSA) is associated with suboptimal disease control and worse chronic severity of asthma. However, little is known about the relations of OSA with acute asthma severity in hospitalized patients.

Objective: To investigate the association of OSA with acute asthma severity. **Methods:** This is a retrospective cohort study using State Inpatient Databases from eight geographically-diverse US states, 2010-2013. The outcomes were markers of acute severity—mechanical ventilation use, hospital length-of-stay (LOS), and inhospital mortality. To determine the association of interest, we fit multivariable logistic regression models adjusting for age, sex, race/ethnicity, primary insurance, household income, patient residence, comorbidities, hospital state, and hospitalization year. We repeated the analysis for children aged 6-17 years.

Results: Among 73,408 adult patients hospitalized for acute asthma, 10.3% had OSA. Coexistent OSA was associated with a significantly higher risk of non-invasive positive pressure ventilation (NIPPV) use (14.9% vs. 3.1%; unadjusted OR 6.48 [95%CI 5.88-

7.13]; adjusted OR 5.20 [95%CI 4.65-5.80]), while coexistent OSA was not significantly associated with the risk of invasive mechanical ventilation use. Patients with OSA had 37% longer hospital LOS (unadjusted incidence rate ratio [IRR] 1.37 [95%CI 1.33-

1.40]); this significant association persisted in the multivariable model (IRR 1.13 [95%CI 1.10-1.17]). The in-hospital mortality did not significantly differ between groups. These findings were consistent in both obesity and non-obesity groups, and in 27,935 children.

Conclusion: Among patients hospitalized for acute asthma, OSA was associated with a higher risk of NIPPV use and longer LOS compared to those without OSA.

1 **Association of obstructive sleep apnea with severity of patients hospitalized for acute**
2 **asthma**

3

4 **INTRODUCTION**

5 Asthma is a common inflammatory disease of the airways, affecting approximately 27
6 million Americans in 2016.¹ Although asthma mortality has declined,² the acute morbidity
7 remains substantial. Indeed, acute asthma accounts for approximately 340,000 hospitalizations in
8 the U.S. each year.³ In parallel, obstructive sleep apnea (OSA) is another common chronic
9 respiratory condition. Recent studies have indicated that OSA affects approximately 20% of the
10 U.S. population⁴ and coexists in 8% to 50% of patients with asthma.^{5,6}

11 Increasing evidence indicates that, among patients with asthma, coexistent OSA is
12 associated with poor disease control.^{4,7-9} For example, observational studies have reported that,
13 compared to the patients without OSA, those with coexistent OSA have a higher Asthma Control
14 Questionnaire score,⁸ more severe daytime and nighttime symptoms,¹⁰ worse quality of life,^{10,11}
15 and more frequent exacerbations.^{7,11} In addition, another study has also reported that the patients
16 with both asthma and OSA have increased healthcare utilization (e.g., higher hospital charges).¹²
17 ^{13,14} While the literature has demonstrated the link between OSA and chronic morbidity of
18 asthma, the relationship between OSA and acute severity measures among patients hospitalized
19 for acute asthma remains to be elucidated. Hospitalized asthma patients are an important
20 population with high morbidity and large healthcare burden.⁶

21 To address this knowledge gap in the literature, we analyzed a large, population-based
22 dataset from eight racially/ethnically- and geographically-diverse U.S. states to investigate the
23 association of coexistent OSA with acute asthma severity. We hypothesized that patients with

24 OSA who were hospitalized for acute asthma have a higher risk of non-invasive or invasive
25 positive pressure ventilation use, longer hospital length-of-stay (LOS), and in-hospital mortality
26 when compared to those without OSA.

27

28 **METHODS**

29 **Study Design and Setting**

30 We conducted a retrospective cohort study using data from the 2010-2013 State Inpatient
31 Databases (SIDs) of eight US states (Arkansas, California, Florida, Iowa, Nebraska, New York,
32 Utah, and Washington). The SID is a component of the Healthcare Cost and Utilization Project
33 (HCUP) sponsored by the Agency for Healthcare and Research Quality. The HCUP data are the
34 largest collection of longitudinal hospital care data in the U.S. with all-payer, encounter-level
35 information. The HCUP SID encompass approximately 97 percent of all U.S. community
36 hospital discharges, and contain all inpatient discharges from short-term, acute-care, non-federal,
37 general, and other specialty hospitals—regardless of payers, source of hospitalization, or
38 disposition—in the participating states. Additional details of the SID may be found elsewhere.¹⁵
39 These eight states were selected for their geographic distribution and high data quality. The
40 institutional review board of Massachusetts General Hospital approved this study.

41 **Study Sample**

42 We identified all unplanned hospitalizations made by patients aged 18-54 years with a
43 primary discharge diagnosis of asthma (*International Classification of Diseases, Ninth Revision,*
44 *Clinical Modification [ICD-9-CM] codes: 493.xx*).¹⁶⁻¹⁸ Then, we further identified patients with
45 OSA by using a concurrent diagnosis of OSA (*ICD-9-CM codes: 327.23, 780.53, and 780.57*) in
46 any diagnosis field, according to prior literature.^{19,20} We also analyzed data focusing on children

47 aged 6-17 years since asthma and OSA are prevalent in this population. The lower cut-off value
48 of age was determined according to the Global Initiative for Asthma (GINA) guidelines since no
49 tests diagnose asthma with certainty in children 5 years and younger.²¹ We included only the first
50 hospitalization for acute asthma for each patient during the study period.

51

52 **Measurements**

53 The SID contains the information on patient demographics (age, sex, and race/ethnicity),
54 primary insurance, estimated household income, urban-rural status, patient comorbidities,
55 hospital state, hospitalization year, *ICD-9-CM* diagnoses, procedures, and disposition. The cut-
56 offs for the estimated income quartile designation were determined using ZIP code-demographic
57 data. The urban-rural status of the patient residence was defined according to the National
58 Center for Health Statistics guidelines.²²

59

60 **Outcomes**

61 The primary outcomes were the use of non-invasive mechanical ventilation (NIPPV;
62 *ICD-9-CM* procedure code 93.90) or invasive positive pressure ventilation (codes 96.04 and
63 96.70-96.72) during the hospitalization, hospital length-of-stay (LOS), and in-hospital
64 mortality.^{17,23}

65

66 **Statistical Analysis**

67 First, we examined the patient characteristics at the hospitalization for acute asthma.
68 Next, to examine the association between OSA and each outcome, we fit unadjusted and
69 multivariable logistic regression models using generalized estimating equations to account for

70 patient clustering within hospitals. In the multivariable models, we adjusted for age (18-39 and
71 40-54 years for adults), sex, race/ethnicity (non-Hispanic white, non-Hispanic black, Hispanic,
72 Asian or Pacific Islander, Native American, and others), primary insurance (Medicare, Medicaid,
73 private, no insurance, and others), quartiles for median household income, patient residence
74 (metropolitan and non-metropolitan residence), 28 Elixhauser comorbidity measures²⁴ as well as
75 arrhythmia,²⁵ hospital state, and hospitalization year, based on biological plausibility and *a priori*
76 knowledge.^{17,18,23,26} For the hospital LOS outcome, we constructed two models—1) logistic
77 regression model using the hospital LOS as a dichotomous variable (LOS ≤ 3 days vs. LOS ≥ 4
78 days based on the median LOS in the data) and 2) negative binomial model fitting the LOS as a
79 count variable.

80 To determine the robustness of our inference, we also performed a series of sensitivity
81 analyses. First, we repeated the analysis with the stratification by the concurrent diagnosis of
82 obesity (*ICD-9-CM* codes: 278.00, 278.01, v85.31-v85.39, and v85.41-85.45) because obesity
83 exists in 70% of patients with OSA.^{17,26} Second, we repeated the analysis with a stratification by
84 age (18-39 vs. 40-54 years) and sex (male vs. female). Third, we used the stabilized inverse
85 probability weighting (IPW) method to estimate the effect of OSA on the outcomes in this
86 observational study.²⁷ Weighting subjects by an inverse probability to have the exposure (OSA)
87 creates a synthetic sample in which the exposure is independent from measured baseline
88 covariates—i.e., in the synthetic sample, OSA and non-OSA patients are exchangeable with
89 regard to the risk factors for the outcomes. Although conventional IPW enables us to obtain
90 unbiased estimates of average effects of OSA on each outcome, patients with a very low or high
91 probability of having the exposure can increase the variability of the estimated effects. In
92 contrast, the stabilized IPW method addresses this problem and directly estimates both the main

93 effect and its variance using conventional regression models.²⁷ All analyses were performed
94 using STATA 14.1 (StataCorp, College Station, TX). All P-values were two-tailed, with $P < 0.05$
95 considered statistically significant.

Journal Pre-proof

96 RESULTS

97 Patient Characteristics

98 During the 4-year study period, we identified 73,408 adult patients hospitalized for acute
99 asthma across the eight U.S. states. Overall, the median age was 44 years (interquartile range
100 [IQR] 33-49 years), 70% were women, and 45% were non-Hispanic white. Of these, 7,564
101 patients (10.3%) had a concurrent OSA. The patients with OSA were older and more likely to be
102 male, non-Hispanic white, and Medicare beneficiaries, compared to those without OSA (all,
103 $P<0.001$; **Table 1**). These patients with OSA were also more likely to have comorbidities, such
104 as hypertension, diabetes, and congestive heart failure (all $P<0.001$).

105

106 OSA and Severity Outcomes

107 **Figure 1** and **Supplemental Table 1** summarize the unadjusted and adjusted associations
108 of OSA with each outcome. Patients with a concurrent diagnosis of OSA had a significantly
109 higher risk of NIPPV use compared to those with non-OSA (14.9% vs. 3.1%; unadjusted OR
110 6.48 [95% CI 5.88-7.13]; adjusted OR 5.20 [95% CI 4.65-5.80]) in the patients with OSA, while
111 there was no significant association of OSA with the risk of invasive mechanical ventilation use.
112 Similarly, the patients with OSA had a higher risk of prolonged hospital LOS (i.e., $LOS \geq 4$ days)
113 compared to those without OSA (66.0% vs. 47.9%; unadjusted OR 2.06 [1.96-2.17]; adjusted OR
114 1.39 [95% CI 1.31-1.48]). Likewise, in the analysis using the hospital LOS as a count variable,
115 the patients with OSA had a 37% longer hospital LOS (unadjusted incidence rate ratio [IRR]
116 1.37; 95% CI 1.33-1.40). This significant association also persisted after adjusting for the
117 potential confounders and patient clustering (adjusted IRR 1.13; 95% CI 1.10-1.17). By contrast,
118 there was no statistically significant difference in in-hospital mortality (0.15% vs. 0.16%;

119 unadjusted OR 0.93 [95%CI 0.49-1.77]; adjusted OR 0.46 [95%CI 0.21-1.01]) between the
120 patients with OSA and those without.

121

122 **OSA and Severity Outcomes in children**

123 The associations between OSA and acute asthma severity persisted in the analysis of
124 27,935 children aged 6-17 years with acute asthma. Overall, 395 (1.4%) had a diagnosis of
125 coexistent OSA. Patient characteristics are shown in **Supplemental Table 2**. Children with OSA
126 were likely to be older and to have public health insurance (Medicaid). Among the children with
127 acute asthma, similar to the findings in adults, coexistent OSA was associated with a
128 significantly higher risk of NIPPV use and longer hospital LOS (both $P < 0.001$; **Supplemental**
129 **Table 3**).

130

131 **Sensitivity Analysis**

132 **Table 2** summarizes the associations between OSA and acute severity of acute asthma,
133 according to obesity status. In this sensitivity analysis, and similar to the main findings,
134 concurrent OSA was associated with a significantly higher risk of NIPPV use both in the non-
135 obesity (adjusted OR 4.98; 95%CI 4.23-5.88) and obesity (adjusted OR 5.49; 95%CI 4.73-6.36)
136 groups. Likewise, OSA was associated with a longer hospital LOS both in the non-obesity
137 (adjusted IRR 1.14; 95%CI 1.08-1.20) and obesity (adjusted IRR 1.14; 95%CI 1.09-1.18)
138 groups. In the stratified analysis by age (**Supplemental Table 4**), the associations between OSA
139 and outcomes were similar to the main findings, while the magnitude of the association with the
140 use was perhaps amplified in the older patients (age 40-54 years). Likewise, in the sensitivity
141 analysis stratified by sex (**Supplemental Table 5**), OSA was associated with a significantly

142 higher risk of NIPPV use and longer hospital LOS in both men and women. Furthermore, all of
143 these associations remained significant in the sensitivity analysis using the stabilized IPW
144 method (**Supplemental Table 6**).

Journal Pre-proof

145 **DISCUSSION**

146 In this population-based study of 73,408 adult patients and 27,935 children hospitalized
147 for acute asthma in eight U.S. states, we found that concurrent OSA was associated with a
148 significantly higher risk of NIPPV use. In addition, these patients with coexistent OSA had an
149 approximately 40% longer hospital LOS compared to those without OSA. By contrast,
150 concurrent OSA and asthma was not associated with significantly higher inpatient mortality. All
151 of these associations persisted after stratifying by obesity status. Furthermore, the observed
152 associations persisted across several different analytic assumptions (i.e., the stratification by age
153 and sex, and analysis using stabilized IPW).

154 The literature has shown that OSA (diagnosed by symptoms or polysomnography) is not
155 only prevalent in patients with asthma^{9,28-31} but also contributes to chronic morbidity of asthma.<sup>5-
156 8,32</sup> For example, in a single-center study of 472 adults with asthma, a higher Sleep Apnea scale
157 of the Sleep Disorders Questionnaire score was associated with a higher risk of poorly-controlled
158 asthma—defined by the Asthma Control Questionnaire score of ≥ 1.5 .⁸ This finding was validated
159 by an analysis of 401 subjects (255 patients with asthma and 146 health controls) who are
160 enrolled in the Severe Asthma Research Program (SARP) II, which also reported the associations
161 with more severe asthma symptoms, more frequent short-acting β -agonist use and healthcare
162 utilization, and worse quality of life.⁷ Furthermore, studies reported that coexistent OSA is
163 associated with higher frequencies of acute asthma.^{13,14} Another study using nationally-
164 representative inpatient data also showed that patients with coexistent asthma and OSA had
165 higher total hospital charges (while the cost information was not available).¹² The present study
166 builds on these prior reports, and extends them by comprehensively demonstrating the relations

167 of OSA with increased severity of acute asthma—i.e., the higher risk of NIPPV use and
168 prolonged hospital LOS—among patients hospitalized for acute asthma.

169 In the current study, comorbid OSA was not significantly associated with the risk of
170 invasive mechanical ventilation use, whereas a previous study using US nationally-representative
171 inpatient data reported increased respiratory therapy including invasive positive pressure
172 ventilation use in asthma patients with OSA.¹² The apparent discrepancy in the results between
173 the earlier and our studies may be attributable to the difference in the definition of outcome
174 measure (i.e., intubation or respiratory therapy). Indeed, the previous study defined “respiratory
175 therapy (intubation and mechanical ventilation)” using *Clinical Classification Software (CCS)*
176 code of 216 in the *primary CCS*-procedure filed, which includes NIPPV use. Therefore, the
177 positive association between OSA and intubation therapy observed in the earlier study was
178 driven, at least partially, by the higher risk of NIPPV use—which is consistent with our findings.

179

180 **Potential Limitations**

181 The current study has several potential limitations. First, as with any study using
182 administrative data, there may been some misclassifications (e.g., underestimation of OSA) in
183 the current study. However, this would have increased the outcome risks preferentially in the
184 non-OSA group, thereby biasing the inferences toward the null. In addition, the HCUP data have
185 been validated against the National Hospital Discharge Survey. Second, the SIDs do not include
186 some of the helpful clinical information on chronic severity measures for asthma (e.g., chronic
187 symptoms, controller use, and pulmonary function) and OSA (e.g., polysomnography,
188 symptoms). Third, as with any observational study, the causal inference might be confounded by
189 unmeasured factors, such as chronic severity of asthma, severity of OSA, and institutional

190 variations in resource use. Yet, the observed associations between OSA and severity of acute
191 asthma remained significant after accounting for at least hospital-level variations. Fourth, our
192 findings are not validated using *ICD-10-CM* codes. However, the use of *ICD-9-CM* codes to
193 identify asthma has high specificity (93%) and negative predictive value (82%) compared with
194 the reference standard using manual chart review by a clinician,³³ supporting the validity of
195 observed relations between two disease conditions (rather than those between *ICD*-coded
196 diagnoses). Finally, while the study sample was comprised of racially/ethnically- and
197 geographically-diverse patients with asthma in the eight U.S. states, our inferences might not be
198 generalizable to patients with less-severe acute asthma (e.g., those who presented to the
199 emergency department without a subsequent hospitalization). Nevertheless, our data remain
200 directly relevant for 340,000 patients hospitalized for acute asthma in the US each year³—a
201 population with high morbidity and large healthcare utilization.

202
203 In summary, by using large population-based data of 73,408 adult patients and 27,935
204 children hospitalized for acute asthma in eight U.S. states, we found that the patients with
205 coexistent OSA had a significantly higher risk of NIPPV use and prolonged hospital LOS
206 compared to those without OSA. These associations persisted after adjusting for potential
207 confounders and across several different analytic assumptions. For clinicians, our findings
208 underscore the importance of accurately identify patients at high risk, such as patients with
209 coexistent OSA and acute asthma. For researchers, our observations should facilitate further
210 investigations into the pathobiological mechanisms that underlie the identified OSA-acute
211 severity association in asthma and encourage the development of targeted prevention and
212 treatment strategies in this clinical population with high morbidity.

213

214 **REFERENCES**

- 215 1. Centers for Disease Control and Prevention. Most Recent Asthma Data. Webpage last
216 updated: May 2018. https://www.cdc.gov/asthma/most_recent_data.htm. Accessed July
217 11, 2019.
- 218 2. American Lung Association Epidemiology and Statistics Unit Research and Health
219 Education Division. Trends in Asthma Morbidity and Mortality. 2012. Accessed July 11,
220 2019.
- 221 3. Agency for Healthcare Research and Quality HCUPnet. <https://hcupnet.ahrq.gov/#setup>.
222 Accessed December 1, 2018.
- 223 4. Peppard PE, Young T, Barnet JH, Palta M, Hagen EW, Hla KM. Increased prevalence of
224 sleep-disordered breathing in adults. *Am J Epidemiol.* 2013;177(9):1006-1014.
- 225 5. Kong DL, Qin Z, Shen H, Jin HY, Wang W, Wang ZF. Association of Obstructive Sleep
226 Apnea with Asthma: A Meta-Analysis. *Sci Rep.* 2017;7(1):4088.
- 227 6. Davies SE, Bishopp A, Wharton S, Turner AM, Mansur AH. The association between
228 asthma and obstructive sleep apnea (OSA): A systematic review. *J Asthma.* 2018:1-12.
- 229 7. Teodorescu M, Broytman O, Curran-Everett D, et al. Obstructive Sleep Apnea Risk,
230 Asthma Burden, and Lower Airway Inflammation in Adults in the Severe Asthma
231 Research Program (SARP) II. *J Allergy Clin Immunol Pract.* 2015;3(4):566-575 e561.
- 232 8. Teodorescu M, Polomis DA, Hall SV, et al. Association of obstructive sleep apnea risk
233 with asthma control in adults. *Chest.* 2010;138(3):543-550.
- 234 9. Julien JY, Martin JG, Ernst P, et al. Prevalence of obstructive sleep apnea-hypopnea in
235 severe versus moderate asthma. *J Allergy Clin Immunol.* 2009;124(2):371-376.

- 236 10. Teodorescu M, Polomis DA, Gangnon RE, et al. Asthma Control and Its Relationship
237 with Obstructive Sleep Apnea (OSA) in Older Adults. *Sleep Disord.* 2013;2013:251567.
- 238 11. ten Brinke A, Sterk PJ, Masclee AA, et al. Risk factors of frequent exacerbations in
239 difficult-to-treat asthma. *Eur Respir J.* 2005;26(5):812-818.
- 240 12. Becerra MB, Becerra BJ, Teodorescu M. Healthcare burden of obstructive sleep apnea
241 and obesity among asthma hospitalizations: Results from the U.S.-based Nationwide
242 Inpatient Sample. *Respir Med.* 2016;117:230-236.
- 243 13. Kim MY, Jo EJ, Kang SY, et al. Obstructive sleep apnea is associated with reduced
244 quality of life in adult patients with asthma. *Ann Allergy Asthma Immunol.*
245 2013;110(4):253-257, 257 e251.
- 246 14. Wang Y, Liu K, Hu K, et al. Impact of obstructive sleep apnea on severe asthma
247 exacerbations. *Sleep Med.* 2016;26:1-5.
- 248 15. Overview of the State Inpatient Databases (SID). Healthcare Cost and Utilization Project.
249 Agency for Healthcare Research and Quality. [http://www.hcup-](http://www.hcup-us.ahrq.gov/sidoverview.jsp)
250 [us.ahrq.gov/sidoverview.jsp](http://www.hcup-us.ahrq.gov/sidoverview.jsp). . Accessed December 1, 2018.
- 251 16. Taille C, Rouvel-Talleg A, Stoica M, et al. Obstructive Sleep Apnoea Modulates Airway
252 Inflammation and Remodelling in Severe Asthma. *PLoS One.* 2016;11(3):e0150042.
- 253 17. Luthe SK, Hirayama A, Goto T, Faridi MK, Camargo CA, Jr., Hasegawa K. Association
254 Between Obesity and Acute Severity Among Patients Hospitalized for Asthma
255 Exacerbation. *J Allergy Clin Immunol Pract.* 2018.
- 256 18. Hasegawa K, Gibo K, Tsugawa Y, Shimada YJ, Camargo CA, Jr. Age-Related
257 Differences in the Rate, Timing, and Diagnosis of 30-Day Readmissions in Hospitalized
258 Adults With Asthma Exacerbation. *Chest.* 2016;149(4):1021-1029.

- 259 19. Louis JM, Mogos MF, Salemi JL, Redline S, Salihu HM. Obstructive sleep apnea and
260 severe maternal-infant morbidity/mortality in the United States, 1998-2009. *Sleep*.
261 2014;37(5):843-849.
- 262 20. Chen Y-H, Kang J-H, Lin C-C, Wang I-T, Keller JJ, Lin H-C. Obstructive sleep apnea
263 and the risk of adverse pregnancy outcomes. *American journal of obstetrics and*
264 *gynecology*. 2012;206(2):136. e131-136. e135.
- 265 21. GINA ASTHMA. [http://ginasthma.org/wp-](http://ginasthma.org/wp-content/uploads/2016/01/GINA_Report_2015_Aug11-1.pdf)
266 [content/uploads/2016/01/GINA_Report_2015_Aug11-1.pdf](http://ginasthma.org/wp-content/uploads/2016/01/GINA_Report_2015_Aug11-1.pdf). Accessed July 11, 2019.
- 267 22. Ray WA, Murray KT, Hall K, Arbogast PG, Stein CM. Azithromycin and the risk of
268 cardiovascular death. *N Engl J Med*. 2012;366(20):1881-1890.
- 269 23. Goto T, Hirayama A, Faridi MK, Camargo CA, Jr., Hasegawa K. Obesity and Severity of
270 Acute Exacerbation of Chronic Obstructive Pulmonary Disease. *Ann Am Thorac Soc*.
271 2018;15(2):184-191.
- 272 24. Elixhauser A, Steiner C, Harris DR, Coffey RM. Comorbidity measures for use with
273 administrative data. *Med Care*. 1998;36(1):8-27.
- 274 25. Thompson NR, Fan Y, Dalton JE, et al. A new Elixhauser-based comorbidity summary
275 measure to predict in-hospital mortality. *Med Care*. 2015;53(4):374-379.
- 276 26. Hirayama A, Goto T, Shimada YJ, Faridi MK, Camargo CA, Jr., Hasegawa K.
277 Association of Obesity With Severity of Heart Failure Exacerbation: A Population-Based
278 Study. *J Am Heart Assoc*. 2018;7(6).
- 279 27. Austin PC, Stuart EA. Moving towards best practice when using inverse probability of
280 treatment weighting (IPTW) using the propensity score to estimate causal treatment
281 effects in observational studies. *Stat Med*. 2015;34(28):3661-3679.

- 282 28. Janson C, Gislason T, Boman G, Hetta J, Roos BE. Sleep disturbances in patients with
283 asthma. *Respir Med.* 1990;84(1):37-42.
- 284 29. Auckley D, Moallem M, Shaman Z, Mustafa M. Findings of a Berlin Questionnaire
285 survey: comparison between patients seen in an asthma clinic versus internal medicine
286 clinic. *Sleep Med.* 2008;9(5):494-499.
- 287 30. Teodorescu M, Consens FB, Bria WF, et al. Correlates of daytime sleepiness in patients
288 with asthma. *Sleep Med.* 2006;7(8):607-613.
- 289 31. Yigla M, Tov N, Solomonov A, Rubin AH, Harlev D. Difficult-to-control asthma and
290 obstructive sleep apnea. *J Asthma.* 2003;40(8):865-871.
- 291 32. Teodorescu M, Polomis DA, Teodorescu MC, et al. Association of obstructive sleep
292 apnea risk or diagnosis with daytime asthma in adults. *J Asthma.* 2012;49(6):620-628.
- 293 33. Wu ST, Sohn S, Ravikumar KE, et al. Automated chart review for asthma cohort
294 identification using natural language processing: an exploratory study. *Ann Allergy*
295 *Asthma Immunol.* 2013;111(5):364-369.
- 296

297 **FIGURE LEGEND**

298 **Figure 1. Unadjusted and adjusted associations between obstructive sleep apnea and acute**
299 **severity of asthma exacerbation**

300

301 Obstructive sleep apnea (OSA) was associated with a significantly higher risk of NIPPV use. The
302 patients with OSA had a 37% longer hospital length-of-stay compared to those without OSA in
303 the unadjusted model. The association remained significant after adjusting for age, sex,
304 race/ethnicity, primary insurance, quartiles for median household income, patient residence, 28
305 Elixhauser comorbidity measures as well as arrhythmia, hospital state, and hospitalization year.

Table 1. Characteristics of patients hospitalized for acute asthma, according to coexistence of obstructive sleep apnea

Characteristics	Obstructive	No	P value
	sleep apnea n=7,564 (10.3%)	obstructive sleep apnea n=65,844 (89.7%)	
Age, median (IQR), year	47 (40-51)	43 (32-49)	<0.001
Female	5,077 (67.1)	46,215 (70.5)	<0.001
Race/ethnicity			<0.001
Non-Hispanic white	3,620 (49.6)	27,799 (44.1)	
Non-Hispanic black	2,165 (29.7)	17,534 (27.8)	
Hispanic	1,166 (16.0)	13,133 (20.9)	
Asian or Pacific Islander	84 (1.2)	1,071 (1.7)	
Native American	42 (0.6)	265 (0.4)	
Others*	218 (3.0)	3,172 (5.0)	
Primary health insurance			<0.001
Medicare	1,802 (23.8)	7,461 (11.3)	
Medicaid	2,506 (33.1)	22,721 (34.5)	
Private	2,274 (30.1)	20,462 (31.1)	
No insurance	597 (7.9)	10,822 (16.4)	
No charge	115 (1.5)	1460 (2.2)	
Others	266 (3.5)	2,882 (4.4)	
Quartiles for median household income			0.02
1 (lowest)	2,846 (39.1)	24,036 (38.7)	
2	1,811 (24.9)	15,320 (24.7)	
3	1,642 (22.5)	13,488 (21.7)	
4 (highest)	986 (13.5)	9,211 (14.8)	
Patient residence			0.70
Metropolitan	6,974 (92.5)	60,726 (92.7)	
Non-metropolitan	590 (7.7)	5,118 (7.3)	
Selected comorbidities†			
Hypertension	4,704 (62.2)	20,463 (31.1)	<0.001
Diabetes	3,041 (40.0)	9,861 (15.0)	<0.001
Congestive heart failure	1,216 (16.1)	2,400 (3.6)	<0.001
Cardiac arrhythmias	931 (12.3)	6,326 (9.6)	<0.001
Renal failure	442 (5.8)	1,249 (1.9)	<0.001

Hospital state			<0.001
Arkansas	215 (2.8)	2,088 (3.2)	
California	1,163 (15.4)	11,035 (16.8)	
Florida	2,753 (36.4)	20,741 (31.5)	
Iowa	211 (2.8)	1,343 (2.0)	
Nebraska	129 (1.7)	938 (1.4)	
New York	2,275 (30.1)	24,807 (37.7)	
Utah	127 (1.7)	991 (1.5)	
Washington	691 (9.1)	3,901 (5.9)	
Hospitalization year			0.77
2010	2,699 (35.7)	23,780 (36.1)	
2011	2,114 (28.0)	18,408 (28.0)	
2012	1,453 (19.2)	12,647 (19.2)	
2013	7,564 (17.2)	11,089 (16.7)	

Data are shown as n (%) unless otherwise specified.

* The other insurance status includes worker's compensation, unreimbursed native health, other miscellaneous.

† Selected from Elixhauser comorbidity

Table 2. Unadjusted and adjusted associations between obstructive sleep apnea and severity of acute asthma, according to obesity

Obesity status and outcomes	Obstructive sleep apnea (95% CI)	No obstructive sleep apnea (95% CI)	Unadjusted association (95% CI)	P value	Adjusted association* (95% CI)	P value
Non-obesity (n=55,307)						
Non-invasive positive pressure ventilation	12.2% (10.9%-13.7%)	3.0% (2.9%-3.2%)	5.14 (4.42-5.98)	<0.001	4.98 (4.23-5.88)	<0.001
Invasive mechanical ventilation	1.8% (1.3%-2.5%)	2.0% (1.8%-2.1%)	0.97 (0.70-1.35)	0.85	1.09 (0.94-2.13)	0.64
Hospital length-of-stay \geq 4 days	61.0% (58.9%-63.1%)	45.9% (45.5%-46.3%)	1.82 (1.66-1.99)	<0.001	1.41 (1.28-1.56)	<0.001
Hospital length-of-stay as a count variable, day, median (IQR)	3 (2-5)	2 (1-4)	1.28 (1.22-1.35)†	<0.001	1.14 (1.08-1.20)†	<0.001
In-hospital mortality	0.14% (0.05%-0.44%)	0.16% (0.13%-0.20%)	0.90 (0.27-2.95)	0.86	0.37 (0.08-1.66)	0.19
Obesity (n=18,101)						
Non-invasive positive pressure ventilation	16.0% (15.0%-17.0%)	2.4% (3.1%-3.8%)	5.91 (5.15-6.79)	<0.001	5.49 (4.73-6.36)	<0.001
Invasive mechanical ventilation	1.7% (1.4%-2.1%)	1.4% (1.2%-1.6%)	1.17 (0.91-1.52)	0.23	0.98 (0.72-1.33)	0.91
Hospital length-of-stay \geq 4 days	67.9% (66.7%-69.2%)	56.1% (55.2%-56.9%)	1.64 (1.53-1.75)	<0.001	1.40 (1.30-1.51)	<0.001
Hospital length-of-stay as a count variable, day, median (IQR)	3 (2-5)	3 (2-4)	1.25 (1.20-1.30)†	<0.001	1.14 (1.09-1.18)†	<0.001
In-hospital mortality	0.15% (0.07%-0.29%)	0.15% (0.10%-0.24%)	0.97 (0.42-2.25)	0.95	0.50 (0.19-1.36)	0.18

Abbreviations: CI, confidence interval; IQR, interquartile range.

Associations are indicated by odds ratio unless otherwise specified.

* Logistic regression model for the binomial outcomes and negative binomial model for the count outcome (hospital length-of-stay), adjusting for age, sex, race/ethnicity, primary insurance, household income, residential status, comorbidities, hospital state, and year.

† Incidence rate ratio.

